

Prenatal Risk Predicts Preschooler Executive Function: A Cascade Model

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Little research has considered whether prenatal experience contributes to executive function (EF) development above and beyond postnatal experience. This study tests direct, mediated, and moderated associations between prenatal risk factors and preschool EF and IQ in a longitudinal sample of 1,292 children from the Family Life Project. A composite of prenatal risk factors (i.e., low birth weight, prematurity, maternal emotional problems, maternal prepregnancy obesity, and obstetric complications) significantly predicted EF and IQ at age 3, above quality of the postnatal environment. This relationship was indirect, mediated through infant general cognitive abilities. Quality of the postnatal home and child-care environments did not moderate the cascade model. These findings highlight the role of prenatal experience as a contributor to individual differences in cognitive development.

Research has increasingly recognized the role of the early caregiving environment in directing the development of executive function (EF), a set of higher order cognitive abilities thought to be necessary for planning and executing goal-directed behavior. Because the specific skills that comprise EF (working memory, inhibitory control, and attention shifting) are rapidly developing between ages 3 and 5 (Garon, Bryson, & Smith, 2008), many of these studies have focused on caregiving's contribution to EF development in the preschool years. For example, high-quality home and child-care environments, marked by the presence of responsive caregivers and provision of learning materials, seem to make a positive contribution to the development of EF in

this time period (Clark et al., 2013; Sarsour et al., 2011). Some evidence suggests that these factors may also serve as a buffer against the negative impacts of poverty and other more distal influences (Berry et al., 2016; Nelson et al., 2015).

In the home environment, children's experiences of sensitive and supportive caregiving in infancy and toddlerhood have specifically been shown to aid in the development of EF (for a review, see Fay-Stammach, Hawes, & Meredith, 2014). Specific parenting practices such as scaffolding (Bernier, Carlson, Deschênes, & Matte-Gagné, 2012; Bernier, Carlson, & Whipple, 2010; Hammond, Müller, Carpendale, Bibok, & Liebermann-Finestone, 2012; Hughes & Ensor, 2009) and attention directing/redirecting (Conway & Stifter, 2012) have also been investigated. This body of literature has concluded that interactions with sensitive caregivers, who provide opportunities and guidance to children in harnessing their emerging executive abilities, are indispensable for supporting optimal EF development.

However, relatively little research has considered whether there are earlier experiences, such as those afforded by the fetal environment, that make an independent contribution to EF development, controlling for postnatal interactions with sensitive caregivers in enriched learning environments. Recent literature has begun to elucidate the role of the prenatal environment in influencing child

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neurocognitive development more generally (Bale et al., 2010; O'Donnell, O'Connor, & Glover, 2009), suggesting that a similar framework could prove useful in the prediction of children's EF trajectories. Theoretically, these extant findings provide support for the prenatal programming hypothesis, whereby prenatal experiences are thought to have long-lasting implications for infant health and development (O'Connor, Heron, Golding, Glover, & the ALSPAC Study Team, 2003). Similar to the prenatal programming hypothesis, the Developmental Origins of Health and Disease (DOHaD; Gillman, 2005) and fetal origins (Barker, 1995) hypotheses emphasize the importance of studying the earliest origins of health and developmental outcomes, by investigating adverse exposures during critical windows of development (including the prenatal period) that may confer increased risk. All three of these hypotheses are similar in their view that prenatal experiences may affect fetal development in ways that predispose individuals to disease and adverse developmental outcomes later in life. Depending on the ontogeny of the impacted systems, these predispositions may be further mediated or moderated by postnatal experiences, but nevertheless provide a starting point from which to understand subsequent child outcomes. This study draws upon this framework to test the precursors and pathways that predict EF in young children.

Prenatal Programming of EF

There is accumulating evidence supporting the prenatal programming of cognitive functioning, and specifically EF. Meta-analyses on outcomes for children born premature or low birth weight (LBW), two purported proxy variables for adverse prenatal exposure, have shown long-term deficits in IQ (Aylward, Pfeiffer, Wright, & Verhulst, 1989) and EF (Aarnoudse-Moens, Weisglas-Kuperus, van Goudoever, & Oosterlaan, 2009) for these children, compared to full term, normal birth weight peers. Prospective, longitudinal studies beginning in pregnancy have begun to elucidate more specific prenatal risk factors that predict impaired EF development.

A set of studies have found that maternal emotional problems, such as depression, state anxiety, and pregnancy-specific anxiety predict worse EF in school-aged children (Buss, Davis, Hobel, & Sandman, 2011; O'Donnell et al., 2017). These findings are often explained in terms of the deleterious effects of high maternal cortisol during pregnancy on the development of the hypothalamus–

pituitary–adrenal axis (O'Connor, Bergman, Sarkar, & Glover, 2013), with these effects possibly operating through epigenetic modification (Oberlander et al., 2008; Sandman, Davis, Buss, & Glynn, 2011). Elevated levels of maternal stress and cortisol have also been shown to alter child neurodevelopment (Sarkar et al., 2014), which would similarly have implications for behavior and cognitive development. A parallel line of work investigating maternal medical problems, such as obesity and obstetric complications, has reached similar conclusions regarding prenatal programming of EF. Two studies found that maternal prepregnancy obesity is associated with poorer child EF, compared to children of normal weight or overweight mothers (Buss et al., 2012; Theresia et al., 2017). One explanation for these findings is that maternal prepregnancy obesity contributes to a state of systemic inflammation, which exerts adverse prenatal programming effects on child neurodevelopment (Van Der Burg et al., 2016). Similarly, maternal obstetric complications (e.g., gestational diabetes, preeclampsia) have been shown to be related to infant neurodevelopmental outcomes, including risk of attention deficit hyperactivity disorder (Seidman et al., 2000; Sprich-Buckminster, Biederman, Milberger, Faraone, & Lehman, 1993). However, these findings have not yet been extended to include EF as a cognitive outcome.

As illustrated by the studies described earlier, much of the research testing the prenatal programming hypothesis has thus far taken a piecemeal approach to identifying risk factors that independently predict EF. That is, single categories of risk factors (e.g., maternal emotional problems) have been studied in isolation, controlling for other possible confounds (e.g., maternal medical problems). While these studies have been successful in pinpointing a group of possible prenatal contributors, correlations among individual risk factors may obscure meaningful prediction of outcomes (Mosteller & Tukey, 1977), thus necessitating the use of a different approach. Cumulative risk approaches have benefits over single indicator models, including greater stability and increased power (Burchinal, Roberts, Hooper, & Zeisel, 2000; Evans, Li, & Whipple, 2013), and have been used extensively to study postnatal risk, including in the current sample (e.g., Burchinal & Willoughby, 2013).

One cumulative approach involves modeling prenatal conditions as a latent variable. Two studies have used this technique, modeling favorable fetal growth conditions as a latent variable with reflective indicators of infant birth weight, birth length,

and gestational age (Bollen, Noble, & Adair, 2013; Camerota & Bollen, 2016). This model has been shown to provide better fit when compared to a model that includes each of the three indicators as independent, observed variables, although how this latent variable predicts child outcomes has not yet been tested.

A second cumulative approach utilizes data reduction techniques (e.g., principal components analysis [PCA], summation of dichotomous risk factors) to create a single indicator of prenatal risk, using multiple risk factors that have previously been studied in isolation (e.g., low birth weight, prematurity, maternal emotional problems, maternal prepregnancy obesity, and obstetric complications). Historically, this approach has been used to categorize LBW or preterm infants into groups of higher or lower medical risk, to better explain heterogeneity in developmental outcomes (e.g., Laucht, Esser, & Schmidt, 1997). More recently, cumulative prenatal risk (i.e., sum of maternal pregnancy complications, infant LBW, and prematurity) has been found to predict poorer social cognition in toddlers (Mark Wade, Madigan, Akbari, & Jenkins, 2015). Similarly, a composite of perinatal complications (e.g., infant Apgar scores less than six, respiratory distress syndrome) was found to predict lower IQ and less mature brain activation in adolescents born preterm (Carmody et al., 2006). Finally, Silveira et al. (2017) recently found that a cumulative prenatal adversity index (e.g., maternal chronic disease, smoking, and mental health during pregnancy, infant prematurity) significantly predicted child neurodevelopmental outcomes in two community birth cohorts, with the cumulative index acting as a stronger predictor than any isolated risk factor. Thus, there is precedence for using cumulative measures of prenatal risk, and some evidence to suggest that these indices are more predictive than individual risk factors. Therefore, applying a cumulative approach has the potential to enhance our understanding of the contributions of prenatal experience to child EF.

A second limitation of previous studies examining prenatal influences on EF is that they have not included comprehensive, objective measurement of the postnatal caregiving environment, beyond proxy variables such as maternal education or socioeconomic status. This omission is problematic for several reasons. For one, the prenatal and postnatal environment are likely correlated, leading to inflated estimates of the effect of the prenatal environment if the postnatal caregiving environment is not also considered. Furthermore, the prenatal and

postnatal environments may interact to predict developmental outcomes. For example, a number of studies have found that sensitive caregiving interacts with various indicators of prenatal risk (e.g., birth weight, prematurity, medical risk) to predict child cognitive outcomes (Camerota, Willoughby, Cox, & Greenberg, 2015; Madigan, Wade, Plamondon, Browne, & Jenkins, 2014; Wade et al., 2015). These findings suggest that the developmental trajectories associated with high prenatal risk are potentially modifiable by high-quality postnatal environments. However, it is currently unknown what factors in the postnatal environment, beyond maternal sensitivity, may have this protective effect, although this knowledge could inform prevention and intervention efforts.

Therefore, there is a need for additional studies that consider both prenatal and postnatal environmental influences on development, in order to elucidate what is specific to the prenatal period, as well as what aspects of the postnatal environment may exacerbate or ameliorate prenatal programming effects.

Role of Foundational Cognitive Abilities

Beyond uncovering the specific contributions of prenatal and postnatal experience to child development, it is also important to understand the earliest cognitive manifestations associated with adverse prenatal conditions, as these inquiries have clear implications for early detection and intervention efforts. It is well-established that child cognition shows at least some stability from infancy through early childhood (Bornstein & Sigman, 1986), with one study showing an indirect pathway from infant habituation time to child full-scale IQ, through the intermediary processes of general cognitive and language abilities (Bornstein et al., 2006). This indirect pathway may continue even further downstream to explain individual differences in adolescent academic achievement (Bornstein, Hahn, & Wolke, 2013). These studies illustrate the concept of a developmental cascade (e.g., Cox, Mills-Koonce, Propper, & Gariépy, 2010), the process by which functioning at one level of behavior (e.g., general cognitive abilities) affects functioning in later-developing, higher level domains (e.g., IQ, achievement).

When considering EF specifically, there are a host of cognitive precursors emerging within the first 3 years of life that may exert cascading influences on EF development in the preschool years and beyond (for a review, see Hendry, Jones, & Charman, 2016). These precursors may include

general cognitive abilities such as information processing and attention. Individual differences in these early developing cognitive domains may be partially responsible for the long-term effects of prenatal conditions on children's higher order cognitive functioning. Indeed, many studies document deficits in general cognitive capacities (e.g., mental development scores, attention) for children born preterm and/or low birth weight (Murray et al., 2014), and among children exposed to high levels of prenatal maternal stress (Huizink, Robles de Medina, Mulder, Visser, & Buitelaar, 2003) and obstetric complications (Brinksma et al., 2017).

Several studies have found support for a cascade model linking foundational cognitive abilities to later differences in EF. Two studies have specifically examined infant information processing as an early cognitive precursor to preschool and child EF. Cuevas and Bell (2014) found that infants who were more efficient at processing a novel stimulus at 5 months (i.e., exhibited shorter median looking time) exhibited better EF performance on a battery of tasks when they were 24, 36, and 48 months. Similarly, Rose, Feldman, and Jankowski (2012) demonstrated that infant processing speed and memory at 7 and 12 months predicted better EF when children were 11 years old. A corollary question pertinent to the current investigation is the extent to which prenatal conditions are indirectly related to later EF through their effect on foundational cognitive abilities.

Only two studies have tested a full cascade model linking the prenatal environment to EF via foundational cognitive ability. Using a small sample of preterm and matched full-term 11-year-olds, researchers found that the relationship between prematurity and poorer EF was partially mediated through children's slower processing speeds (Rose, Feldman, & Jankowski, 2011). Furthermore, the negative effects of prematurity on academic achievement were entirely mediated through processing speed and EF. However, as all cognitive measures were administered concurrently in this study, it is unclear how the development of these processes unfolds longitudinally in children who experience different prenatal conditions. A second longitudinal study found that the relationship between birth weight and EF at age 4 was indirect, operating through language ability at age 3 (Wade, Browne, Madigan, Plamondon, & Jenkins, 2014). However, it remains unclear whether these same relationships would be observed using a more comprehensive measure of prenatal risk. This issue constitutes the main impetus for the current study.

The Current Study

The overarching goal of the current investigation is to test a cascade model by which variations in prenatal conditions are linked to preschool cognitive outcomes (i.e., EF and IQ) through the mediating mechanism of infant general cognitive ability. In testing this model, we pose three research questions. First, to what extent do prenatal conditions predict child EF and IQ at age 3, controlling for quality of the postnatal environment? Although previous literature has explored the role of individual prenatal risk factors, we adopt a cumulative risk approach that tests whether variations in a wide range of variables indicative of prenatal risk (i.e., low birth weight, prematurity, maternal emotional problems, maternal prepregnancy obesity, and obstetric complications) predict preschoolers' cognition, in a large, socioeconomically, and ethnically diverse sample. Furthermore, we include outcome measures of both EF and IQ in this investigation, in order to determine whether prenatal programming effects are specific to EF, or whether they impact child cognitive development more broadly. Given that previous research on EF has not tended to include additional cognitive measures, such as IQ scores, it is unknown whether the magnitude of prenatal programming effects will be similar for both cognitive outcomes.

Our second question is whether the relationship between prenatal risk and preschool cognitive outcomes operates indirectly through infant general cognitive ability. Consistent with previous studies finding support for cascading effects of early cognitive deficits (e.g., Rose, Feldman, Jankowski, & Rossen, 2008; Rose et al., 2011), we predict that the relationship will be largely indirect, with more risky prenatal environments predicting worse general cognition in infancy, which will negatively predict downstream EF. Because our mediator of infant cognitive abilities includes general cognitive processes, rather than specific precursors of EF, we predict that this cascade model will similarly predict deficits in IQ.

Finally, we test whether the relationship between prenatal risk and preschool cognitive outcomes is moderated by the postnatal environment, specifically observed caregiver responsiveness and ratings of child-care quality. Previous research in this sample has demonstrated that mothers' (Blair, Raver, & Berry, 2014; Blair et al., 2011) and fathers' (Towe-Goodman et al., 2014) sensitive parenting from 6 to 36 months predicts child EF from ages 3 through 5. There is also evidence in this sample that sensitive

parenting (Camerota et al., 2015) and high-quality child care (Berry et al., 2016) buffer children from delays in EF given adverse prenatal or postnatal experiences. We expand these extant findings by simultaneously considering the postnatal caregiving environment in two settings (i.e., home and child care). We predict that higher quality home and child-care environments will buffer children from the negative effects of risky prenatal environments, by moderating the links from prenatal risk to infant general cognitive ability, and from infant general cognitive ability to EF and IQ.

Method

Participants

These data come from the Family Life Project (FLP), a longitudinal study of children living in primarily rural, low-income areas of the United States. Specifically, three counties in North Carolina and three counties in Pennsylvania were targeted. Participants were recruited during a 1-year period from September 2002 through September 2003, at the time that they gave birth to a child. Low-income and African American families were oversampled in order to address study goals. Detailed sampling and recruitment information exist elsewhere (Burchinal, Vernon-Feagans, Cox, & Key Family Life Project Investigators, 2008). Of the 5,471 families who gave birth during the recruitment period, 72% met eligibility criteria. Women were excluded if they did not live in a selected county, spoke a primary language other than English, or intended to move away from the target area within the next 3 years. The final sample consisted of 1,292 families who completed the 2-month home visit, at which point they were considered officially enrolled in the study. In the final sample, 51% of children were male and 41% were African American.

Procedures

These analyses use data collected from home visits when children were approximately 2, 6, 15, 24, and 36 months of age. Home visits lasted approximately 2–3 hr and consisted of parent (e.g., semistructured interviews), child (e.g., cognitive assessments), and dyadic (e.g., parent–child interaction tasks) activities. Pertinent to the current investigation, mothers retrospectively reported on pregnancy experiences during the 2-month home visit. At 6, 15, 24, and 36 months, infant and their caregivers engaged in a semistructured free play

activity which was recorded for later coding of parental warm sensitivity and harsh intrusiveness. At each of these time points, research assistants also rated the quality of the child-care environments, for children who attended child care outside of the home. At 15 months, children's general cognitive ability was assessed. At 36 months, children completed additional cognitive assessments including tasks to assess EF and IQ. All study protocols were approved by the institutional review boards at [blinded], and appropriate consent and assent were obtained from all participants.

Measures

Prenatal Risk Factors

At 2 months, mothers responded to the pregnancy and delivery module of the Missouri Assessment of Genetics Interview for Children (MAGIC; Todd, Joyner, Heath, Neuman, & Reich, 2003), an interview surveying maternal experiences during pregnancy, including emotional problems and medical complications during pregnancy and delivery. The MAGIC module has been previously validated, with one study finding that mothers demonstrate good reliability and stability in reporting pre/perinatal events (Reich, Todd, Joyner, Neuman, & Heath, 2003).

Three yes/no questions were used to determine mothers' emotional problems during pregnancy. These included, "Were you happy when you found out you were pregnant with [target child]?", "Did you experience emotional problems, such as depression or anxiety, which upset you a lot, or for which you sought treatment or counseling?", and "Did you experience serious family problems which were upsetting to you?" The first question was reversed scored and added to the latter two questions to yield a sum score of emotional problems (range = 0–3). Nine medical problems during pregnancy were assessed, including heavy bleeding, excessive nausea or vomiting, weight loss over 10 pounds, water retention, convulsions, accidents requiring medical care, hypertension, infection, and other illnesses requiring medical care. The sum of all pregnancy complications was used as a composite of medical risk (range = 0–9).

Mothers self-reported their prepregnancy weight in pounds, as well as their current height in feet and inches. Using this information, we calculated mothers' prepregnancy body mass index (BMI) and used this to determine whether she was obese (BMI \geq 30) before her pregnancy with the target

child. Mothers also reported her infant's birth weight in pounds and ounces, which was converted into grams and used to categorize infants as low birth weight (<2,500 g). Finally, mothers reported infant due date and birth date, which was used to calculate infant gestational age and subsequently, prematurity (< 37 weeks).

Descriptive statistics on all prenatal risk indicators are presented in Table 1. We utilized data reduction techniques to decrease the number of independent variables, and facilitate planned mediation and moderation analyses. Because of our interest in capturing variability among our entire set of prenatal risk factors, we conducted a PCA and retained the first principal component as our measure of prenatal risk. This approach allowed us to retain information about individual predictors while limiting the number of correlated independent variables included in analysis. We preferred this approach to a risk index (i.e., a count score) because it did not require arbitrary dichotomization of continuous variables, thus retaining the most

information among its underlying set of indicators (Burchinal et al., 2000).

PCA indicated that all five prenatal variables (i.e., maternal emotional problems, pregnancy complications, maternal prepregnancy obesity, low birth weight, and prematurity) loaded positively onto a single factor (loadings = .28–.76). This factor explained 31% of the variance in the underlying items. A single factor score was output for each individual and is subsequently referred to as prenatal risk, with higher values indicating more risk. Additional information about this factor score, including its distribution and correlation with individual indicators, is presented in Appendix S1.

Postnatal Caregiving Quality

Mother–child interactions were coded from video recordings of semistructured play activities that lasted approximately 10 min. At 6 and 15 months, the dyad engaged in a free play activity. Parents were given a standardized set of toys (e.g., stacking rings, shape sorter) and instructed to play with their child as they normally would if they had some free time during the day. At 24 and 36 months, the dyads engaged in a puzzle task, in which experimenters provided a set of developmentally appropriate puzzles of increasing difficulty for children to complete. Mothers were instructed to provide any assistance that they thought was necessary. In line with extant coding schemes (Cox & Crnic, 2002), these interactions were coded for the constructs of Sensitivity, Intrusiveness, Detachment, Stimulation of Cognitive Development, Positive Regard, Negative Regard, and Animation. Global ratings of parents' behaviors were made on a 1–5 scale at 6 and 15 months, and a 1–7 scale at 24 and 36 months, with values ranging from 1 (*not at all characteristic*) to 7 (*highly characteristic*) of the parent, based on the following definitions.

Sensitivity measured the degree to which the parent was aware of and responsive to the child's signals and cues, and engaged in synchronous interaction with the infant. Intrusiveness was defined by the parent controlling the interaction in a parent-centered way, such as inappropriately fast pacing or overstimulation. Detachment describes the degree to which a parent was emotionally distant or unaware of the child's signals. Stimulation is defined by the parent engaging in age-appropriate behaviors that foster the child's cognitive development, such as labeling objects. Positive Regard rates the quantity and intensity of the parent's positive feelings toward the child, while Negative

Table 1
Descriptive Statistics for Prenatal Risk Variables

Variable name	N	M (SD)/%
<i>Maternal emotional problems</i>	1,287	0.47 (0.72)
Were you happy when found out you were pregnant?	1,282	82
Did you experience emotional problems, such as depression or anxiety, for which you sought treatment or counseling?	1,287	13
Did you experience serious family problems which were upsetting to you?	1,287	16
<i>Pregnancy complications</i>	1,287	1.45 (1.31)
Heavy bleeding	1,287	4.5
Excessive nausea or vomiting	1,287	28
Weight loss > 10 pounds	1,286	18
Infection	1,287	27
Hypertension	1,287	17
Water retention	1,287	27
Convulsions	1,287	0.4
Accident requiring medical care	1,287	8.6
Other illness requiring medical care	1,287	13
<i>Maternal prepregnancy obesity</i>	1,238	24
Prepregnancy body mass index	1,238	26.4 (7.16)
<i>Infant low birth weight</i>	1,287	8.0
Birth weight (g)	1,287	3276 (585.9)
<i>Infant premature</i>	1,274	8.0
Gestational age (weeks)	1,274	39.12 (1.91)

The italic values represent the composite or coded values that went into the principal components analysis.

Regard rates the quantity and intensity of negative feelings toward the child. Finally, Animation captures the degree of parental energy and animation during the interaction, which could be vocal, physical, or affective.

Based on exploratory factor analysis conducted with oblique rotation (i.e., Promax), which suggested two underlying factors, we created composite scores of warm sensitivity and harsh intrusiveness by taking the mean across multiple parenting scales (Mills-Koonce et al., 2011; Vernon-Feagans, Cox, & FLP Key Investigators, 2013). Warm sensitivity comprised of scores from the Sensitivity, Detachment (reverse coded), Stimulation of Cognitive Development, Positive Regard, and Animation scales, while harsh intrusiveness comprised of scores from the Intrusiveness and Negative Regard scales. We used the composite scores of warm sensitivity and harsh intrusiveness in our analyses, as these scores capture both positive and negative dimensions of parenting. This composite score approach has been widely used in previous analyses with FLP data (Blair et al., 2014). High scores on the sensitivity composite indexed parental behavior that was warm, responsive, stimulating, and child-centered, whereas high scores on the intrusiveness composite indexed parental behavior that was harsh, negative, and parent-focused.

All coders were trained and certified by one master coder. In addition, reliability checks were completed on approximately 30% of tapes to ensure intraclass correlations (ICC) between all pairs of coders exceeded .80 (ICC = .80–.98 for all subscales and composite scores). To test the impact of timing of parental input, and to facilitate planned moderation analyses, we averaged warm sensitivity and harsh intrusiveness scores across 6–15 and 24–36 months to represent caregiver quality during infancy and toddlerhood, respectively.

Postnatal Child-care Quality

Traditional child-care quality measures (e.g., Early Childhood Environment Rating Scale; Harms & Clifford, 1980) were unavailable in these data. Therefore, consistent with previous studies (e.g., Berry et al., 2016), we included independent ratings of the quality of the child-care environment using the Home Observation Measure of the Environment scale (HOME; Caldwell & Bradley, 1984). Quality of the child-care environment was only rated for children who attended child care outside of the home ($N = 454, 505, 427, 455$ at 6, 15, 24, and 35 months, respectively). Child-care arrangements could

include formal (e.g., center-based) or informal settings (e.g., family day care). At each time point, trained research assistants observed child-care settings for approximately 20 min for the purpose of completing the HOME scale.

The infant/toddler version of the measure was used at 6, 15, and 24 months, whereas the preschool version was used at 36 months. All three subscales (Responsivity, Acceptance, and Learning Materials) were used at all time points, but the number of items per subscale varied based on the version. The infant/toddler version contained 11 items for Responsivity, 8 items for Acceptance, and 9 items for Learning Materials. The preschool version contained 7 items for Responsivity, 4 items for Acceptance, and 11 items for Learning Materials. We summed items from all subscales to create a single index of child-care quality at each time point (range = 0–28 at 6, 15, and 24 months; range = 0–22 at 35 months). Internal consistency of all items was adequate at each of the four time points ($\alpha = .58-.69$). We averaged together ratings of child-care quality for 6 and 15 months (infancy), and 24 and 35 months (toddlerhood) to create two summary indicators of the postnatal child-care environment.

Child Cognitive Outcomes

General cognitive ability. Child developmental status was assessed at 15 months using the Bayley Scales of Infant Development, 2nd ed. (BSID-II; Bayley, 1993). The BSID-II is a widely used measure of cognitive development for children in the first 2 years of life, and measures abilities such as memory, problem solving, and language. The Mental Development Index (MDI), a norm-referenced standard score ($M = 100, SD = 15$), was used as our measure of general cognitive ability.

Executive function. Child EF was assessed at 36 months using a battery of tasks presented in an open spiral bound flip-book format, with pages measuring 8×14 in. Each page presented stimuli to the child on one page and scripted instructions for the research assistant on the other. For each task, children first had to pass a set of training trials, assessing their comprehension of task constructs and procedures, before continuing on to the test trials. The battery of EF tasks included two measures of working memory (Working Memory Span and Pick the Picture Game), three measures of inhibitory control (Silly Sounds Stroop, Spatial Conflict, and Animal Go/No-Go), and one measure of attention shifting (Something's the Same). Details of these individual tasks, as well as psychometric properties of the battery, are explained elsewhere

(Willoughby, Blair, Wirth, & Greenberg, 2010). We therefore omit full task descriptions here. Consistent with the analytic approach previously used in this sample (Willoughby, Wirth, & Blair, 2012), item response theory models were used to create expected a posteriori (EAP) scores for each task at each assessment. Individual EAP scores were averaged and used as an index of child EF ability.

Full-scale IQ. At 36 months, the Block Design and Receptive Vocabulary subscales of the Wechsler Preschool and Primary Scales of Intelligence, 3rd ed. (Wechsler, 2002) were used to assess child full-scale IQ. This two subscale short-form demonstrates the highest correlation ($r = .88$) with full-scale IQ scores (Silverstein, 1975).

Demographic Covariates

At the 2-month home visit, primary caregivers reported on their years of education and their child's gender. Household income-to-needs ratio (INR) was calculated at the 6-, 15-, 24-, and 35-month data collection visits. Aggregated household income from all sources (e.g., earnings, child support) was divided by the U.S. poverty threshold for the year (adjusted for family size and household composition) to create an INR for each family at each time point. Mirroring the approach described above, we averaged together INR at 6 and 15 months and at 24 and 35 months to represent household income in infancy and toddlerhood, respectively.

Missing Data

Of the 1,292 primary caregivers enrolled in this study, 1,287 were biological mothers who could report on their pregnancies at the 2-month home visit. Of these 1,287 biological mothers, 49 (3.8%) were missing prepregnancy obesity data and 9 (0.7%) were missing infant due date data, meaning that prematurity could not be determined. Concerning postnatal data, 200 (16%) children were missing MDI data, 319 (25%) were missing EF data, and 238 (18%) were missing IQ data. There were 87 (6.8%) and 162 (13%) dyads missing ratings of caregiving quality in infancy and toddlerhood, respectively.

Because quality of the child-care environment was only collected for children attending child care outside the home, there is a larger amount of missing data for this variable in infancy ($N = 665$; 52%) and toddlerhood ($N = 701$; 54%). Children who did not have child-care data in infancy were more likely to be White, $\chi^2(1) = 46.03$, $p < .001$, but did not differ based on child gender or poverty status at

recruitment ($p > .05$). Furthermore, children with child-care data in infancy tended to have higher household INR, $t(1,174) = -5.30$, $p < .001$, higher maternal education, $t(1,285) = -2.83$, $p = .005$, and higher levels of maternal intrusiveness, $t(1,203) = -3.04$, $p = .002$, but did not differ on prenatal risk, maternal sensitivity, general cognitive ability, EF, or IQ ($p > .05$). Similar results were found for children with and without child-care data in toddlerhood. Since differences in missing data were accounted for by variables included as covariates in the model, the assumption that these data were conditionally random, or missing at random (MAR), was supported. There was no reason to believe that child-care data were missing because of the level of child-care quality itself, which would render the data not MAR. Thus, we used full-information maximum likelihood for all analyses (Schafer & Graham, 2002). A total of 343 (27%) of dyads had complete data.

Analytic Strategy

To address our first substantive aim, we used the prenatal risk factor score derived above in a total-effects model, to test whether prenatal risk predicted child EF and IQ above and beyond covariates and the postnatal caregiving and child-care environments. Because we hypothesized a cascading relationship between prenatal risk and preschool cognitive outcomes, we next tested a mediation model where we added an indirect path from prenatal risk to preschool EF and IQ through infant general cognitive ability (i.e., MDI score). Using this cascade model, we compared the significance and magnitude of the direct and indirect effects of prenatal risk on EF and IQ. Finally, we tested whether the cascade model described in Step 2 was moderated by the postnatal caregiving environment. We accomplished this by adding six interaction terms to the indirect effects model (Prenatal Risk \times Infant Caregiver Sensitivity, Prenatal Risk \times Infant Caregiver Intrusiveness, Prenatal Risk \times Infant Child-care Quality, MDI \times Toddler Caregiver Sensitivity, MDI \times Toddler Caregiver Intrusiveness, and MDI \times Toddler Child-care Quality), where the first three interaction terms predicted MDI and the latter three predicted EF and IQ. Nonsignificant interaction terms were trimmed and the final model was interpreted.

Descriptive statistics and path models addressing all substantive aims were estimated in Mplus version 8 (Muthén & Muthén, 2017). Path models allowed us to estimate a single model predicting

multiple dependent variables (i.e., EF and IQ). All results accounted for the complex sampling design by adjusting for population weight and stratification variables, and all dependent variables (i.e., MDI, EF, and IQ) were regressed on all covariates (e.g., INR). Additionally, all models accounted for possible relationships among exogenous variables (e.g., stability in parenting from infancy to toddlerhood, relationship between poverty and child-care quality) by allowing these variables to covary. Indirect effects were calculated as the products of coefficients linking the focal predictor to the outcome via the mediating variable(s) and the significance of these effects were determined using the delta method (Mackinnon, 2008).

Results

Preliminary Analyses

Weighted descriptive statistics, including means, standard deviations, and correlations for all study variables are displayed in Table 2. From this table, it is clear that prenatal risk was significantly, though modestly, correlated with child MDI ($r = -.17, p < .001$), EF ($r = -.10, p < .01$), and IQ ($r = -.14, p < .001$). Infant MDI was positively correlated with preschool EF ($r = .23, p < .001$) and IQ ($r = .39, p < .001$). Child male gender was negatively correlated with all cognitive outcomes of interest ($r = -.10$ to $-.16, p < .01$). Household INR was positively related to all cognitive outcomes in infancy and toddlerhood ($r = .17$ to $.36, p < .001$). Infant MDI was positively correlated with caregiver sensitivity ($r = .22, p < .001$) and child-care quality ($r = .18, p < .001$) and negatively correlated with caregiver intrusiveness ($r = -.16, p < .001$) in infancy. Similarly, preschoolers' IQ was positively correlated with caregiver sensitivity ($r = .34$ to $.41, p < .001$) and child-care quality ($r = .21$ to $.24, p < .001$) and negatively correlated with caregiver intrusiveness ($r = -.26$ to $-.45, p < .001$) in both infancy and toddlerhood. Preschoolers' EF was significantly correlated with caregiver sensitivity ($r = .27$ to $.31, p < .001$) and intrusiveness ($r = -.22$ to $-.30, p < .001$) in infancy and toddlerhood, but was not correlated with the quality of the child-care environment at either time point ($r = .10, p > .07$).

Total Effects Model

Our first path model included a direct effect of prenatal risk on EF and IQ, controlling for demographic covariates and caregiver and child-care

quality in infancy and toddlerhood (see Figure 1). The model was fully saturated. There were no direct effects of prenatal risk on either EF ($\beta = -.04, p = .24$) or IQ ($\beta = -.06, p = .13$). Caregiver sensitivity ($\beta = .12, p = .02$), intrusiveness ($\beta = -.25, p < .001$), and child-care quality ($\beta = .14, p < .001$) in toddlerhood significantly predicted child IQ, whereas only caregiver sensitivity ($\beta = .11, p = .04$) and intrusiveness ($\beta = -.14, p < .01$) in toddlerhood significantly predicted child EF (Figure 1).

Mediated (Cascade) Model

We next added an indirect path from prenatal risk to EF and IQ through child MDI. An important change in this model is that parenting and child-care quality in infancy were included as predictors of MDI, while parenting and child-care quality in toddlerhood were included as predictors of EF and IQ (see Figure 2). Model fit was adequate; $\chi^2(12) = 51.65, p < .001$, comparative fit index (CFI) = .93, root mean square error of approximation (RMSEA) = .05 (90% RMSEA CI [.037, .065]). In this model, there was not a significant direct effect of prenatal risk on EF ($\beta = -.03, p = .43$) or IQ ($\beta = -.03, p = .41$). Prenatal risk did significantly predict child MDI ($\beta = -.14, p < .001$), which in turn predicted both EF ($\beta = .11, p < .01$) and IQ ($\beta = .22, p < .001$). The indirect path from prenatal risk to EF ($\beta = -.02, p = .02$; 95% CI [-.028, -.003]) and IQ ($\beta = -.03, p < .001$; 95% CI [-.045, -.014]) through MDI was significant. Caregiver sensitivity ($\beta = .12, p = .001$) and child-care quality ($\beta = .16, p < .001$) in infancy significantly predicted MDI. Caregiver sensitivity ($\beta = .11, p < .01$), intrusiveness ($\beta = -.22, p < .001$), and child-care quality ($\beta = .15, p < .001$) in toddlerhood significantly predicted IQ, whereas only caregiver sensitivity ($\beta = .15, p = .001$) and intrusiveness ($\beta = -.14, p < .01$) in toddlerhood significantly predicted child EF.

Interactive Effects Model

In a final model, we added six interaction terms to test whether the indirect effects of prenatal risk on EF and IQ through MDI were moderated by the quality of the postnatal caregiving and child-care environments. This model fit the data well; $\chi^2(21) = 52.78, p < .001$, CFI = .95, RMSEA = .03 (90% RMSEA CI [.023, .046]). However, none of the added interaction terms (i.e., Prenatal Risk \times Infant Caregiver Sensitivity, Prenatal Risk \times Infant Caregiver Intrusiveness, Prenatal Risk \times Infant Child-care Quality, MDI \times Toddler Caregiver Sensitivity,

Table 2
Weighted Correlations and Descriptive Statistics for Study Variables

	1	2	3	4	5	6	7	8	9	10	11	12	13	14
1. Prenatal risk	—													
2. Bayley MDI (15 months)	-.17	—												
3. EF (36 months)	-.10	.23	—											
4. IQ (36 months)	-.14	.39	.40	—										
5. Male	-.00 [±]	-.12	-.10	-.16	—									
6. Maternal education (years)	-.16	.20	.25	.36	.00 [±]	—								
7. Caregiver sensitivity (infant)	-.11	.22	.27	.34	-.02 [±]	.49	—							
8. Caregiver intrusiveness (infant)	.07	-.16	-.22	-.26	.05 [±]	-.35	-.29	—						
9. Child-care quality (infant)	.01 [±]	.18	.10 [±]	.21	.02 [±]	.14	.19	-.23	—					
10. Caregiver sensitivity (toddler)	-.16	.30	.31	.41	-.01 [±]	.52	.69	-.32	.20	—				
11. Caregiver intrusiveness (toddler)	.19	-.31	-.30	-.45	.10	-.39	-.37	.45	-.21	-.57	—			
12. Child-care quality (toddler)	.06 [±]	.13	.10 [±]	.24	.02 [±]	.17	.16	-.06 [±]	.25	.13	-.10	—		
13. Income-to-needs ratio (infant)	-.12	.17	.22	.35	.05 [±]	.58	.39	-.30	.22	.41	-.30	.22	—	
14. Income-to-needs ratio (toddler)	-.13	.20	.22	.36	.04 [±]	.59	.39	-.27	.21	.42	-.31	.23	.87	—
<i>n</i>	1,227	1,087	968	1,049	1,287	1,287	1,200	1,200	622	1,125	1,125	586	1,172	1,147
<i>M</i>	-.02	97.80	-.45	96.90	.52	13.02	2.98	2.23	25.32	3.02	2.21	20.99	2.23	2.18
<i>SD</i>	1.00	10.36	.54	16.72	2.21	2.21	.71	.59	2.30	.69	.71	3.83	1.82	1.66

Note. Sample sizes for correlations range from 448 to 1,287, given variation among participants' study completion. "Infant" measures are averages taken from 6- and 15-month observations. "Toddler" measures are averages taken from 24- and 35-month observations. All correlations are significant at $p < .05$, except where marked ([±] indicates $p > .05$). MDI = Mental Development Index; EF = executive function.

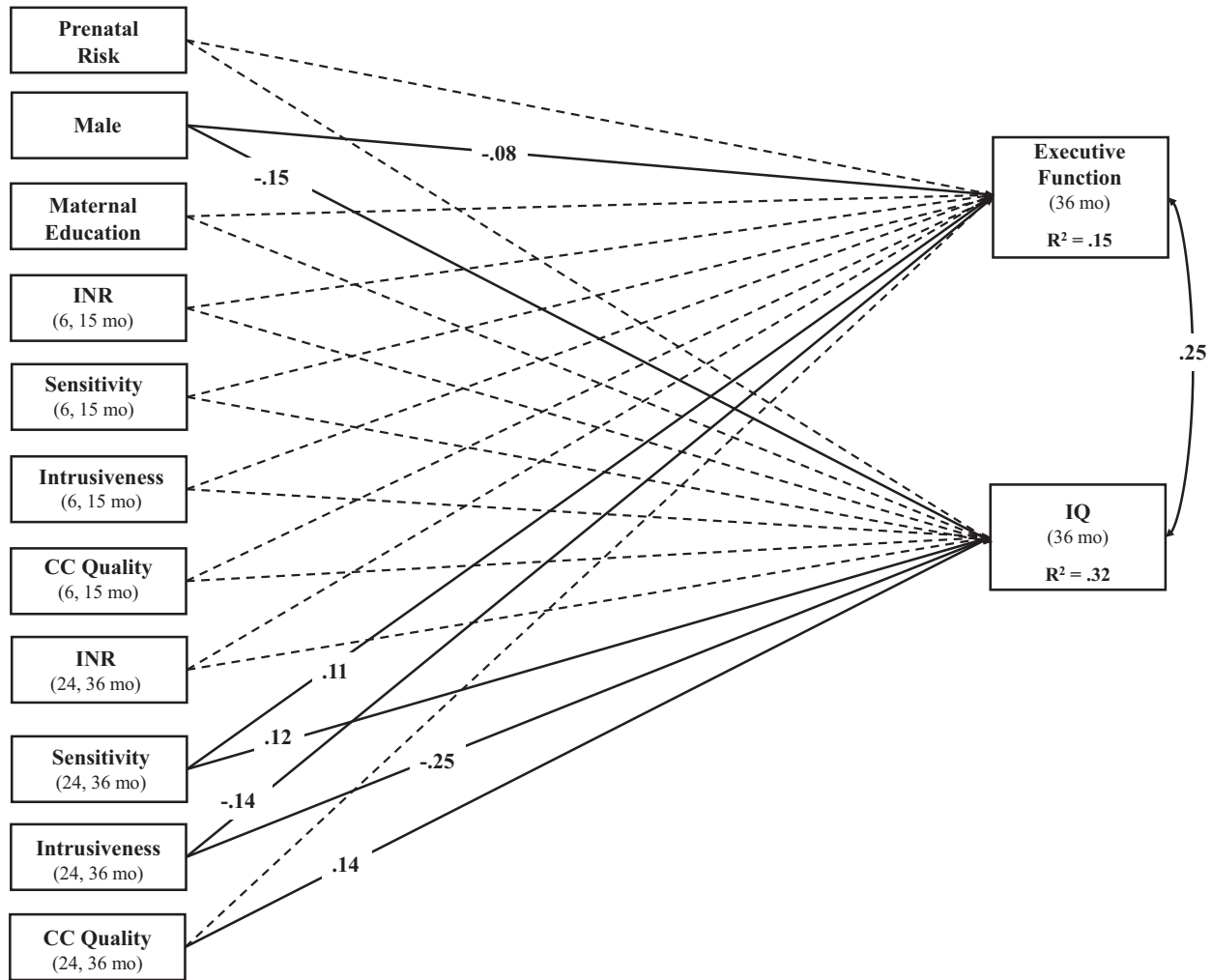


Figure 1. Path model illustrating total effects of prenatal and postnatal factors on EF and IQ at age 3 years. Single-headed arrows represent regression paths, while curved, double-headed arrows represent correlations. Standardized parameter estimates are presented for all significant paths ($p < .05$) using solid lines, whereas non-significant paths are presented using dotted lines. All exogenous variables were allowed to covary. INR = income-to-needs ratio; CC = childcare.

MDI \times Toddler Caregiver Intrusiveness, and MDI \times Toddler Child-care Quality) were significant predictors.

In light of possible concerns regarding multicollinearity among multiple interaction terms, we also tested each moderator individually. That is, we re-estimated six cascade models that each included a single interaction term (e.g., Prenatal Risk \times Infant Child-care Quality). None of these models yielded significant interaction terms, supporting our conclusions from the model that included all six interaction terms simultaneously.

Post Hoc Analyses

Given the absence of interactive effects, we tested the alternative hypothesis that parenting and child-

care quality might mediate the cascade from prenatal risk to general cognitive ability to EF and IQ. That is, we added paths from prenatal risk to parental sensitivity, intrusiveness, and child-care quality in infancy to general cognitive ability, and from general cognitive ability to parental sensitivity, intrusiveness, and child-care quality in toddlerhood to both EF and IQ. This model fit the data well, $\chi^2(9) = 12.55, p = .18, CFI = .99, RMSEA = .02$ (90% RMSEA CI [.000, .038]). In this model, the indirect paths from prenatal risk to MDI through maternal sensitivity ($\beta = -.004, p = .32$), intrusiveness ($\beta = .000, p = .76$), and child-care quality ($\beta = .006, p = .24$) in infancy were all nonsignificant. However, there were significant indirect paths from infant general cognitive ability to parental sensitivity in toddlerhood to EF ($\beta = .02, p = .009$) and IQ

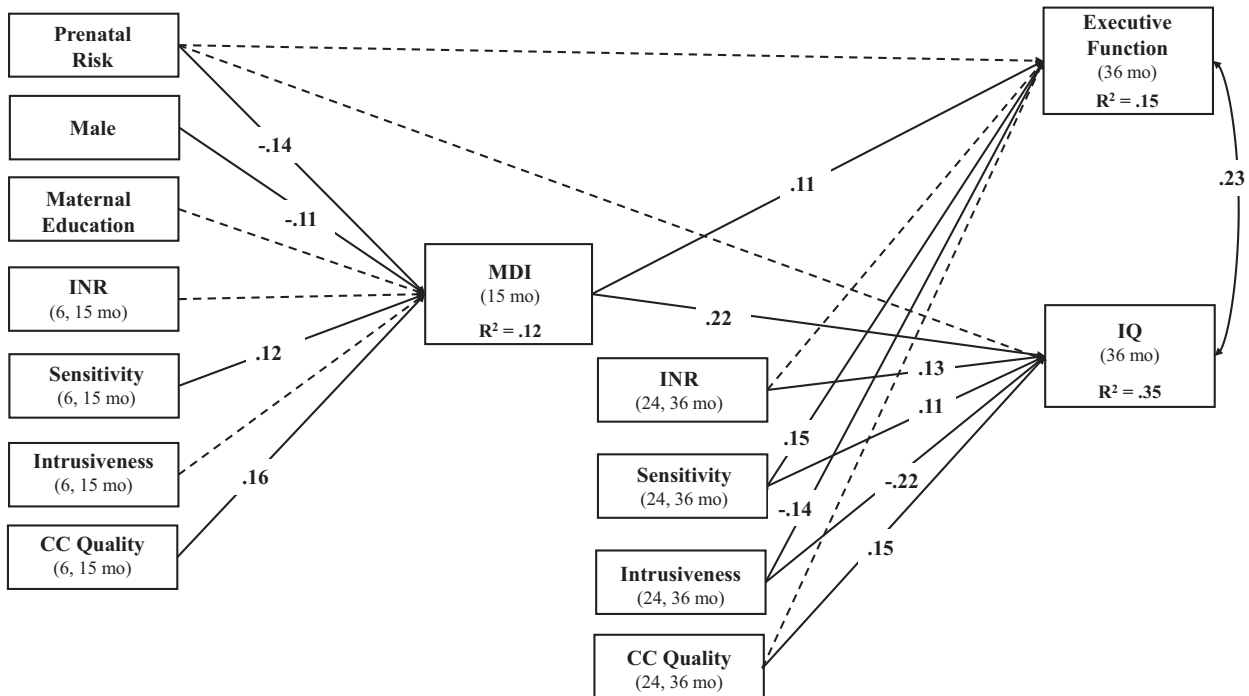


Figure 2. Path model illustrating a cascade of effects from prenatal risk to EF and IQ at age 3 years via general cognitive ability (Bayley Mental Developmental Index [MDI]) at 15 months. Single-headed arrows represent regression paths, while curved, double-headed arrows represent correlations. Standardized parameter estimates are presented for all significant paths ($p < .05$) using solid lines. Non-significant paths are presented using dotted lines. For simplicity, paths from covariates (male, maternal education) to EF and IQ are not pictured. Additionally, all exogenous variables were allowed to covary. INR = income-to-needs ratio; CC = childcare.

($\beta = .01$, $p = .02$). Similarly, the indirect paths from general cognitive ability to parental intrusiveness in toddlerhood to EF ($\beta = .02$, $p = .008$) and IQ ($\beta = .03$, $p < .001$) were significant. The full indirect paths from prenatal risk to infant general cognitive ability to parental sensitivity ($\beta = -.002$, $p = .02$) and intrusiveness ($\beta = -.003$, $p = .02$) to child EF were significant. Similarly, there was a significant indirect path from prenatal risk to infant general cognitive ability to parental sensitivity ($\beta = -.002$, $p = .03$) and intrusiveness ($\beta = -.005$, $p = .003$) predicting child IQ.

Finally, given that infant MDI scores were not adjusted for prematurity, we conducted a sensitivity analysis excluding MDI data from all children born prematurely (< 37 weeks; $n = 81$). That magnitude of all effects remained largely unchanged, although coefficient standard errors increased with the exclusion of these scores, likely due to the decreased sample size. This change resulted in the indirect effect from prenatal risk to EF through general cognitive ability becoming marginally significant ($\beta = -.014$, $SE = .008$, $p = .07$; 95% CI $[-.030, .001]$), compared to the original model ($\beta = -.015$, $SE = .006$, $p = .02$; 95% CI $[-.028, -.003]$). In light

of these relatively minor changes to the size of model coefficients, we concluded that unadjusted MDI scores were unlikely to be driving the effects reported in this manuscript.

Discussion

This study tested the independent and interactive effects of prenatal and postnatal experience on preschoolers' cognitive outcomes. We found that a composite of prenatal risk factors (i.e., low birth weight, prematurity, maternal emotional problems, maternal prepregnancy obesity, and obstetric complications) indirectly predicted child EF and IQ at age 3, above and beyond quality of home and child-care environments in infancy and toddlerhood. This is the first study of which we are aware that has found an independent effect of the prenatal environment on child EF while statistically controlling for the quality of the postnatal environment across several settings. Moreover, we found evidence for a cascade model linking prenatal risk to preschoolers' cognitive outcomes via general cognitive abilities in infancy. Contrary to our hypotheses,

we did not find evidence that the quality of the postnatal home and caregiving environments moderated the relationships among prenatal risk, general cognitive ability, and EF/IQ. Our findings demonstrate support for the DOHaD and prenatal programming hypotheses and suggest the importance of including measures of prenatal conditions in studies of child development.

Environmental contributions to children's cognitive development, and in particular EF, have long been acknowledged in the literature. Thus far, the factors considered have primarily fallen under the domain of postnatal environmental characteristics, including poverty (Blair & Raver, 2012), parenting (Fay-Stammach et al., 2014), and early education (Yoshikawa et al., 2013). These studies have not simultaneously considered the role of the prenatal environment. Likewise, the small but growing set of studies examining prenatal risk and childhood EF have tended to test only single indicators of prenatal risk and have not included comprehensive, objective assessments of the quality of the postnatal environment (Buss et al., 2011; Phua, Rifkin-Graboi, Saw, Meaney, & Qiu, 2012; Theresia et al., 2017; Wade & Jenkins, 2016). Therefore, we currently know very little about the independent contributions of prenatal and postnatal experience to the development of EF. The present study, indicating a small but significant effect of prenatal risk above and beyond the quality of the postnatal home and caregiving environments, corroborate and expand upon these previous findings. To the extent that prenatal and postnatal environmental quality are correlated with one another as well as with child EF, our inclusion of both indicators in the same model serves as a more stringent test of the hypothesis that prenatal experience shapes EF development.

Results of our mediation analyses provide support for a cascade model where the relationship between prenatal experience and preschool EF operates indirectly through infant's fundamental cognitive abilities. In other words, exposure to adverse prenatal environments may initiate a cascade of developmental difficulties, due to their more proximal influence on foundational cognitive abilities (Rose et al., 2008). Our findings are consistent with developmental theories that suggest continuity in cognitive development starting in infancy (Bornstein et al., 2006), and more specifically, with studies that trace the origins of individual differences in EF to earlier differences in fundamental cognitive processes such as attention, language, and information processing (Cuevas & Bell, 2014; Rose

et al., 2012; M. Wade et al., 2014). Because these general cognitive processes are measurable beginning in infancy, they provide an early identification and intervention point for children who may be at risk for later EF difficulties (Hendry et al., 2016), whether due to adverse prenatal experience or other factors.

However, one implication of these findings is that prenatal risk is not likely to influence EF alone. Rather, as observed in this study, our cascade model explained the relationship between prenatal risk and both EF and IQ. It therefore remains to be seen whether there are developmental processes that link prenatal risk to EF specifically, or whether EF is just one of many cognitive abilities that are impacted by exposure to adverse prenatal environments. The inclusion of more discrete cognitive assessments in infancy may have improved our ability to test specific mechanisms by which prenatal risk relates to EF and IQ.

Recent research has elucidated several mechanisms that may be responsible for the long-lasting impact of prenatal experience on child cognitive development. Multiple studies have demonstrated alterations in neural structure and function for children born low birth weight (for a review, see Jobe, 2010) or exposed to elevated levels of maternal stress (Sarkar et al., 2014). In a separate set of studies, degree of neural abnormality was found to predict the extent of cognitive deficits in LBW and preterm infants (Lowe et al., 2011; Woodward, Clark, Bora, & Inder, 2012; Woodward, Clark, Pritchard, Anderson, & Inder, 2011). Therefore, exposure to risk factors in the fetal environment may alter trajectories of cognitive development through their impact on neural development, a process that is rapidly occurring during the prenatal period (Tau & Peterson, 2010).

A second compelling hypothesis concerns the role of maternal-placental-fetal neuroendocrine and immune systems as mediators of prenatal environmental influences on child development. Several reviews have postulated the ways in which the developing fetus acquires and integrates information about its environment using these distinct biological systems, the same systems that are involved in adaptation to stress and neurocognitive development later in the postnatal period (Wadhwa, 2005; Wadhwa, Buss, Entringer, & Swanson, 2010). Specifically, elevated fetal exposure to cortisol (O'Connor et al., 2013; Sandman et al., 2011) and corticotrophin-releasing hormone (Sandman, Wadhwa, Chicz-DeMet, Porto, & Garite, 1999) may be both the result of stressful pregnancy experiences

(e.g., obstetric complications, emotional problems), and the common mechanism by which these diverse experiences exert programming effects, possibly by initiating epigenetic changes in the fetus (Oberlander et al., 2008). Relatedly, shared genetics between mothers and offspring may contribute to both higher risk pregnancies and delays or deficits in child cognitive development (Leve et al., 2012).

Finally, behavioral differences in neonates exposed to higher levels of prenatal risk may interfere with the ability of caregivers to provide the type of sensitive responsiveness in the postnatal period that we know is important for cognitive development (e.g., Mills-Koonce et al., 2015). Studies using within-family designs have found evidence for this mechanism, reporting that siblings exposed to higher prenatal risk (e.g., lower birth weight) receive less sensitive caregiving, on average, compared to lower risk (e.g., higher birth weight) siblings (Browne et al., 2018). The current study found some support for this mechanism in our post hoc analyses, which indicated that deficits in infant general cognitive ability, which are partially due to variations in exposure to prenatal risk, predicted decreased maternal sensitivity and increased maternal intrusiveness, which in turn predicted poorer EF and IQ at preschool. As our study was not well placed to test all possible competing mechanisms, it is unclear how biological or genetic mechanisms contribute in addition to parenting.

Contrary to our hypotheses, we did not find evidence that prenatal and postnatal experience interacted to predict child cognitive outcomes at 15 months or 3 years. These findings suggest that high-quality caregiving and child-care experiences are equally beneficial to children regardless of their prenatal experiences. On one hand, these findings are promising, as they suggest that all children have the potential to reap the benefits of supportive postnatal environments, regardless of level of prenatal risk (Austin et al., 2017). However, in an alternative model, we found that parenting quality partially mediated the cascade from prenatal risk to general cognitive ability to EF and IQ. Therefore, while supportive postnatal environments do not seem to buffer children from cognitive deficits resulting from increased prenatal risk, they may represent one pathway by which prenatal risk exerts enduring consequences for child development. Indeed, these intriguing findings suggest that it is not exposure to prenatal risk per se that results in suboptimal parenting. Rather, it is the subtle deficits in general cognitive ability that may result from such exposures that evoke decreased

sensitivity and increased sensitivity in parents. If replicated, these findings have clear implications for prevention and intervention efforts, as promoting early cognitive development in children exposed to elevated levels of prenatal risk may disrupt the cascade leading to poor cognitive development in preschool and beyond.

Our findings are bolstered by the use of a large, longitudinal sample that contained repeated and objective measurement of key constructs, such as postnatal environmental quality and child cognition. Furthermore, we overcome limitations of convenience sampling through our complex sampling design, which allows us to generalize our findings back to the counties from which our data were drawn. In spite of these strengths, an obvious limitation of this study is our reliance on retrospective measurement of the prenatal environment. While previous studies have found maternal report of prepregnancy weight, obstetric complications, birth weight, and gestational age to be consistent with clinical records (Lederman & Paxton, 1998; Sou, Chen, Hsieh, & Jeng, 2006), there is also the possibility that mothers' retrospective report of specific emotional or medical problems may be incomplete or subject to recall bias. Thus, future studies should attempt to replicate the current findings using prospectively obtained data on salient aspects of the prenatal environment. Furthermore, although we attempted to capture a broad range of normative prenatal experiences, rather than specific, known toxic exposures (e.g., tobacco or alcohol use), there are other experiences that we were not able to capture given the current data. For example, maternal diet (e.g., Monk, Georgieff, & Osterholm, 2013), psychosocial stress (e.g., Laplante et al., 2004), and inflammation (e.g., Bastek, Weber, McShea, Ryan, & Elovitz, 2014) during pregnancy have all been linked to subsequent child outcomes. Among the variables we did include, low birth weight, prematurity, and pregnancy complications loaded most strongly on our factor score, while maternal emotional problems and prepregnancy obesity loaded less strongly. This observation may influence the interpretation of our study findings, and also suggests the need for further evaluation of single- and multidimensional models of prenatal risk. Similarly, although a cumulative risk approach provides important preliminary evidence of the relationship between prenatal experience and child cognitive development, this approach precludes us from identifying specific mechanistic pathways or specific periods of influence (e.g., influences from preconception vs. prenatal vs. perinatal periods).

Clearly, additional research is needed in order to identify quantifiable aspects of the prenatal environment, as well as uncover mechanisms linking prenatal experience to postnatal development.

In sum, our findings demonstrate that increased prenatal risk is indirectly related to poorer EF and IQ during the preschool years via infant general cognitive ability. These findings provide support for DOHaD, prenatal programming, and fetal origins hypotheses first proposed in the medical literature. While some efforts have been made to apply these theories to developmental psychology (e.g., Pluess & Belsky, 2011), we hope that the current findings provide an impetus for others to consider the contributions of prenatal experience as an early driver of individual differences in cognitive processes.

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Supporting Information

Additional supporting information may be found in the online version of this article at the publisher's website:

Appendix S1. Additional descriptive statistics for prenatal risk score.